

## Bayesian Nonparametric Statistics for Fostering Innovation and Discovery in Biomedical Research

#### Yanxun Xu

Department of Applied Math and Statistics Mathematics Institute for Data Science Division of Biostatistics and Bioinformatics The Sidney Kimmel Comprehensive Cancer Center Johns Hopkins University

April, BNP 2022

## **Clinical Trials**



## **Clinical Trials**





Clinicatesterday

Biomarker-based Trials











Clinicatesterday

Biomarker-based Trials











## **Population**

Clinicatesterday

Biomarker-based Trials











## **Population**

Clinicatesterbay













## **Observational Data**

## **Population**

Clinicatesterbay





## 





## **Observational Data**





**Population** 

Clinicatesterday



## Biomarker-based Trials







# <text><text><text><text><text><text><text><text><text>

## **Observational Data**





## **Population**

Clinicatesterbay



## Biomarker-based Trials







# <text><text><text><text><text><text><text><text><text><text>

## **Observational Data**





## Population

Subgroup

Personalization

#### **Rely on incorrect assumptions**

#### **Rely on incorrect assumptions**

• Normality (e.g., t test for sample size calculation in clinical trials).

#### **Rely on incorrect assumptions**

- Normality (e.g., t test for sample size calculation in clinical trials).
- Proportional hazard (e.g., survival data analysis).

#### **Rely on incorrect assumptions**

- Normality (e.g., t test for sample size calculation in clinical trials).
- Proportional hazard (e.g., survival data analysis).
- Strong parametric assumption (e.g., Poisson for count data).

#### **Rely on incorrect assumptions**

- Normality (e.g., t test for sample size calculation in clinical trials).
- Proportional hazard (e.g., survival data analysis).
- Strong parametric assumption (e.g., Poisson for count data).

#### **Oversimplify framework**

#### **Rely on incorrect assumptions**

- Normality (e.g., t test for sample size calculation in clinical trials).
- Proportional hazard (e.g., survival data analysis).
- Strong parametric assumption (e.g., Poisson for count data).

#### **Oversimplify framework**

• Homogeneous population.

#### **Rely on incorrect assumptions**

- Normality (e.g., t test for sample size calculation in clinical trials).
- Proportional hazard (e.g., survival data analysis).
- Strong parametric assumption (e.g., Poisson for count data).

#### **Oversimplify framework**

- Homogeneous population.
- Compare survival outcomes only based on firstling therapy, ignoring the subsequent savage treatments.

• Bayesian

## Bayesian

 $\mathbb{P}(\text{parameters}|\text{data}) \propto \mathbb{P}(\text{data}|\text{parameters})\mathbb{P}(\text{parameters})$ 

## Bayesian

 $\mathbb{P}(\text{parameters}|\text{data}) \propto \mathbb{P}(\text{data}|\text{parameters})\mathbb{P}(\text{parameters})$ 

• BNP: Bayesian models that are not parametric (unbounded/growing/infinite number of parameters)

## Bayesian

 $\mathbb{P}(\text{parameters}|\text{data}) \propto \mathbb{P}(\text{data}|\text{parameters})\mathbb{P}(\text{parameters})$ 

• BNP: Bayesian models that are not parametric (unbounded/growing/infinite number of parameters)

## Bayesian

 $\mathbb{P}(\text{parameters}|\text{data}) \propto \mathbb{P}(\text{data}|\text{parameters})\mathbb{P}(\text{parameters})$ 

• BNP: Bayesian models that are not parametric (unbounded/growing/infinite number of parameters)

#### **Advantages:**

• Uncertainty quantification (e.g., small sample size, decision making)

## Bayesian

 $\mathbb{P}(\text{parameters}|\text{data}) \propto \mathbb{P}(\text{data}|\text{parameters})\mathbb{P}(\text{parameters})$ 

• BNP: Bayesian models that are not parametric (unbounded/growing/infinite number of parameters)

- Uncertainty quantification (e.g., small sample size, decision making)
- Nonparametric

## Bayesian

 $\mathbb{P}(\text{parameters}|\text{data}) \propto \mathbb{P}(\text{data}|\text{parameters})\mathbb{P}(\text{parameters})$ 

• BNP: Bayesian models that are not parametric (unbounded/growing/infinite number of parameters)

- Uncertainty quantification (e.g., small sample size, decision making)
- Nonparametric
- Wide support

## Bayesian

 $\mathbb{P}(\text{parameters}|\text{data}) \propto \mathbb{P}(\text{data}|\text{parameters})\mathbb{P}(\text{parameters})$ 

• BNP: Bayesian models that are not parametric (unbounded/growing/infinite number of parameters)

- Uncertainty quantification (e.g., small sample size, decision making)
- Nonparametric
- Wide support
- Flexible

#### **Trial designs**

## **Trial designs**



#### **Patients clustering**





CTNNB1lbeta-Catenin-R-V CTNNA1lalpha-Catenin-M-V PTENIPTEN-R-V IGF1RIIGF-1R-beta-R-C ARIAR-R-V SRCISrc\_pY527-R-V MAPK1 MAPK3IMAPK pT202 Y204-R-V SHC1IShc pY317-R-NA ACACAIACC1-R-C ACACA ACACBIACC\_pS79-R-V CCNE1lCyclin\_E1-M-V BCL2L1IBcl-xL-R-V RAD51IRad51-M-C CDKN1Alp21-R-C FOXO3IFOXO3a-R-C MRE11AlMre11-R-C STMN1/Stathmin-R-V

## **Trial designs**



#### **Patients clustering**



#### **Continuous treatments**



## **Trial designs**



#### Genetic signatures Continuous treatments



## **Trial designs**

## **Patients clustering**



## Reinforcement learning



## Genetic signatures Continuous treatments



# Outline

## • Part 1: Monday

- **Density estimation** for efficient clinical trial designs
- Regression for precision dosing

- Part 2: Wednesday
  - **Clustering** for subgroup finding
  - Latent feature models for tumor heterogeneity

- Part 3: Friday
  - Estimating treatment effects from observational data

# Outline

## • Part 1: Monday

- **Density estimation** for efficient clinical trial designs
- Regression for precision dosing

## • Part 2: Wednesday

- Clustering for subgroup finding
- Latent feature models for tumor heterogeneity

## • Part 3: Friday

• Estimating treatment effects from observational data

# **Motivating Trial**

#### Pulmonary resection


#### Pulmonary resection



Intraoperative air leaks (IALs) occur in 48 to 75% of patients.

### Pulmonary resection



Intraoperative air leaks (IALs) occur in 48 to 75% of patients.

•Routine use: sutures and stapling devices

### Pulmonary resection



Intraoperative air leaks (IALs) occur in 48 to 75% of patients.

- •Routine use: sutures and stapling devices
- •Risks: postoperative pain, infections, longer hospitalization, economic, etc

### Pulmonary resection



Intraoperative air leaks (IALs) occur in 48 to 75% of patients.

- •Routine use: sutures and stapling devices
- •Risks: postoperative pain, infections, longer hospitalization, economic, etc
- •New way: Progel (liquid sealants)

**Goal:** design a randomized trial comparing

Progel versus standard care!

Goal: design a randomized trial comparing

Progel versus standard care!



Frequency

Goal: design a randomized trial comparing

Progel versus standard care!



Air Leak Length

10

ω

Frequency

**Goal:** design a randomized trial comparing

Progel versus standard care!

- mean=8, sd=8.76
- a two sample one-sided 0.05-level t test with power 0.8 to detect a 25% drop, would require n = 476.



10

ω

ဖ

requency

Goal: design a randomized trial comparing

Progel versus standard care!



- a two sample one-sided 0.05-level t test with power 0.8 to detect a 25% drop, would require n = 476.
- log scale mean = 1.61 and sd = 0.97, it would require n = 280.



Goal: design a randomized trial comparing

Progel versus standard care!



Standard parametric model is INAPPROPRIATE!!

• Free of air leaks immediately following surgery is possible.

- Free of air leaks immediately following surgery is possible.
- Consider the distribution of the time to resolve an air leak  $G_j$  as a mixture:

$$G_j = \nu_j \delta_0 + (1 - \nu_j) M_j, j = 0, 1.$$

- Free of air leaks immediately following surgery is possible.
- Consider the distribution of the time to resolve an air leak  $G_j$  as a mixture:

$$G_j = \nu_j \delta_0 + (1 - \nu_j) M_j, j = 0, 1.$$

•  $M_1$  is a left-shifted version of  $M_0$ .

- Free of air leaks immediately following surgery is possible.
- Consider the distribution of the time to resolve an air leak  $G_j$  as a mixture:

$$G_j = \nu_j \delta_0 + (1 - \nu_j) M_j, j = 0, 1.$$

- $M_1$  is a left-shifted version of  $M_0$ .
  - Progel is inert, it cannot react chemically with the patient's lung tissue

- Free of air leaks immediately following surgery is possible.
- Consider the distribution of the time to resolve an air leak  $G_j$  as a mixture:

$$G_j = \nu_j \delta_0 + (1 - \nu_j) M_j, j = 0, 1.$$

- $M_1$  is a left-shifted version of  $M_0$ .
  - Progel is inert, it cannot react chemically with the patient's lung tissue
  - not a potential source of infection

- Free of air leaks immediately following surgery is possible.
- Consider the distribution of the time to resolve an air leak  $G_j$  as a mixture:

$$G_j = \nu_j \delta_0 + (1 - \nu_j) M_j, j = 0, 1.$$

- $M_1$  is a left-shifted version of  $M_0$ .
  - Progel is inert, it cannot react chemically with the patient's lung tissue
  - not a potential source of infection
  - does not slow down the healing process

- Free of air leaks immediately following surgery is possible.
- Consider the distribution of the time to resolve an air leak  $G_j$  as a mixture:

$$G_j = \nu_j \delta_0 + (1 - \nu_j) M_j, j = 0, 1.$$

- $M_1$  is a left-shifted version of  $M_0$ .
  - Progel is inert, it cannot react chemically with the patient's lung tissue
  - not a potential source of infection
  - does not slow down the healing process
  - does not contribute to air leak formation

# Stochastic ordering on $G_1$ and $G_0$ : Progel may be better, but not worse.

Stochastic ordering on  $G_1$  and  $G_0$ : Progel may be better, but not worse.



*T*: the time (in days) to resolve an air leaks. Define  $Y = \log(T + 1)$ .

*T*: the time (in days) to resolve an air leaks. Define  $Y = \log(T + 1)$ .

$$\bar{U}_j = \int u(Y)G_j(dY), \ j = 0, 1$$

*T*: the time (in days) to resolve an air leaks. Define  $Y = \log(T + 1)$ .

$$\bar{U}_{j} = \int u(Y)G_{j}(dY), \ j = 0, 1$$

$$\bigcup Utility$$

*T*: the time (in days) to resolve an air leaks. Define  $Y = \log(T + 1)$ .

$$\bar{U}_{j} = \int u(Y) G_{j}(dY), \ j = 0, 1$$

$$\bigcup Utility$$

# Trial Design: $Pr(\bar{U}_1 > \bar{U}_0 + \epsilon \mid \text{Data})$

A formal utility elicitation with our clinical collaborator. Both medical and economic:

• the most desirable resolution time is T = 0 (free of air leaks immediately after surgery, although this ideal outcome is almost never seen with standard care).

- the most desirable resolution time is T = 0 (free of air leaks immediately after surgery, although this ideal outcome is almost never seen with standard care).
- early  $(1 \le T \le 5)$  resolution of air leaks is very desirable and therefore the interval [1, 5] received a relatively high utility.

- the most desirable resolution time is T = 0 (free of air leaks immediately after surgery, although this ideal outcome is almost never seen with standard care).
- early  $(1 \le T \le 5)$  resolution of air leaks is very desirable and therefore the interval [1, 5] received a relatively high utility.
- the utilities drop off steeply for later resolution times (T > 5).

- the most desirable resolution time is T = 0 (free of air leaks immediately after surgery, although this ideal outcome is almost never seen with standard care).
- early  $(1 \le T \le 5)$  resolution of air leaks is very desirable and therefore the interval [1, 5] received a relatively high utility.
- the utilities drop off steeply for later resolution times (T > 5).

T (days)	0	5	10	15	20	25	30	35	$\geq$ 40
Utility	100	50	10	6	5	4	3	2	0

### Main Idea

 $\bar{U}_j = \int u(Y)G_j(dY), \ j = 0, 1$ 



### Main Idea

$$\bar{U}_j = \int u(Y) G_j(dY), \ j = 0, 1$$



### Main Idea

$$\bar{U}_j = \int u(Y) G_j(dY), \ j = 0, 1$$



The most commonly used BNP prior p(F) for a random probability measure (Ferguson, 1973):

$$F \sim DP(\alpha, F_0)$$
 with  $F(y) = \sum_{h=1}^{\infty} w_h \delta_{m_h}(y)$ .

The most commonly used BNP prior p(F) for a random probability measure (Ferguson, 1973):

$$F \sim \mathrm{DP}(\alpha, F_0)$$
 with  $F(y) = \sum_{h=1}^{\infty} w_h \delta_{m_h}(y)$ .

• Locations:  $m_h \sim F_0$ , i.i.d.

The most commonly used BNP prior p(F) for a random probability measure (Ferguson, 1973):

$$F \sim DP(\alpha, F_0)$$
 with  $F(y) = \sum_{h=1}^{\infty} w_h \delta_{m_h}(y)$ .

• Locations:  $m_h \sim F_0$ , i.i.d.

• Weights: 
$$w_h = v_h \prod_{l < h} (1 - v_l)$$
 with  $v_h \sim \text{beta}(1, \alpha)$ .

The most commonly used BNP prior p(F) for a random probability measure (Ferguson, 1973):

$$F \sim DP(\alpha, F_0)$$
 with  $F(y) = \sum_{h=1}^{\infty} w_h \delta_{m_h}(y)$ .

• Locations:  $m_h \sim F_0$ , i.i.d.

• Weights: 
$$w_h = v_h \prod_{l < h} (1 - v_l)$$
 with  $v_h \sim \text{beta}(1, \alpha)$ .

### **Stick-breaking**
# **Dirichlet Process (DP)**

#### **Parameters:**

- base measure:  $F_0(A) = E\{F(A)\}$
- total mass  $\alpha$ :  $F(A) \sim \text{beta}(\alpha F_0(A), \alpha F_0(A^c))$ .

# **Dirichlet Process (DP)**

#### **Parameters:**

- base measure:  $F_0(A) = E\{F(A)\}$
- total mass  $\alpha$ :  $F(A) \sim \text{beta}(\alpha F_0(A), \alpha F_0(A^c))$ .

#### **Posterior Inference:**

- Likelihood: p(y | F) = F
- Prior:  $F \sim DP(\alpha, F_0)$
- Posterior:  $p(F \mid y) = DP(\alpha + 1, F^*)$  with  $F^* = \frac{\alpha F_o + \delta_y}{\alpha + 1}.$

# **DP Mixtures**

- DP random measure: discrete F is not appropriate for many problems
- **DP mixture (DPM):** convolution of discrete *F* with (continuous) kernel, e.g., normal (Escobar & West, 1995)

# **DP Mixtures**

- DP random measure: discrete F is not appropriate for many problems
- **DP mixture (DPM):** convolution of discrete *F* with (continuous) kernel, e.g., normal (Escobar & West, 1995)

$$G(y) = \int N(y \mid \theta, \sigma^2) dF(\theta), F \sim DP$$
$$= \sum_{h=1}^{\infty} w_h N(m_h, \sigma^2)$$

# **DP Mixtures**

- DP random measure: discrete F is not appropriate for many problems
- **DP mixture (DPM):** convolution of discrete *F* with (continuous) kernel, e.g., normal (Escobar & West, 1995)

$$G(y) = \int N(y \mid \theta, \sigma^2) dF(\theta), F \sim DP$$
$$= \sum_{h=1}^{\infty} w_h N(m_h, \sigma^2)$$



 $G_{j} = \nu_{j0}\delta_{0} + \sum_{jh}^{\infty} \nu_{jh} N(\theta_{jh}, \sigma^{2})$ *h*=1











(a)  $G_0$  and  $G_1$ 



$$G_{j} = \nu_{j0}\delta_{0} + \sum_{h=1}^{\infty} \nu_{jh} N(\theta_{jh}, \sigma^{2})$$
  
=  $\nu_{j0}\delta_{0} + (1 - \nu_{j0}) \sum_{h=1}^{\infty} w_{h} N(\theta_{jh}, \sigma^{2})$ 

$$= \nu_{j0} \delta_0 + (1 - \nu_{j0}) M_j$$

 $G_0$  and  $G_1$ 





Further prior beliefs can be added.



**Trial Design:**  $Pr(\bar{U}_1 > \bar{U}_0 + \epsilon \mid \text{Data})$ 

Xu et al., 2017

### Simulation Study: Setup

$$Y_{ji} \sim 0.8N(1.5, 0.3^2) + 0.2N(3, 0.3^2), j = 0$$
  
 $Y_{ji} \sim 0.8N(1, 0.3^2) + 0.2N(2.5, 0.3^2), j = 1$ 

**Treatment Group** 



# **Simulation Study: Results**



**Treatment Group** 

• BNP utility-based trial designs.

- BNP utility-based trial designs.
- The utility function is only meaningful if the probability model allows learning about detailed features of the event time distribution, and the nonparametric model is only needed when the decision hinges on such details.

- BNP utility-based trial designs.
- The utility function is only meaningful if the probability model allows learning about detailed features of the event time distribution, and the nonparametric model is only needed when the decision hinges on such details.
- Novelties

- BNP utility-based trial designs.
- The utility function is only meaningful if the probability model allows learning about detailed features of the event time distribution, and the nonparametric model is only needed when the decision hinges on such details.
- Novelties
  - a small-scale trial design (n=48).

- BNP utility-based trial designs.
- The utility function is only meaningful if the probability model allows learning about detailed features of the event time distribution, and the nonparametric model is only needed when the decision hinges on such details.
- Novelties
  - a small-scale trial design (n=48).
  - a convincing case for the need of a full probabilistic description of uncertainties on random probability measures

# Outline

### • Part 1: Monday

- Density estimation for efficient clinical trial designs
- Regression for precision dosing

### • Part 2: Wednesday

- Clustering for subgroup finding
- Latent feature models for tumor heterogeneity

### • Part 3: Friday

• Estimating treatment effects from observational data

Allogeneic Stem Cell Transplantation (alloSCT) is an aggressive therapy for various hematologic diseases, such as lymphocytic leukemia and non-Hodgkins lymphoma.



Allogeneic Stem Cell Transplantation (alloSCT) is an aggressive therapy for various hematologic diseases, such as lymphocytic leukemia and non-Hodgkins lymphoma.



#### Busulfan



Systemic busulfan exposure, characterized by area under the plasma concentration curve, AUC = the delivered dose, is strongly associated with clinical outcome.

- AUC too high  $\Rightarrow$  High risks of severe toxicity.
- AUC too low  $\Rightarrow$  High risks of graft failure and disease recurrence.

Systemic busulfan exposure, characterized by area under the plasma concentration curve, AUC = the delivered dose, is strongly associated with clinical outcome.

- AUC too high  $\Rightarrow$  High risks of severe toxicity.
- AUC too low  $\Rightarrow$  High risks of graft failure and disease recurrence.
- Earlier practice: busulfan orally, resulting in 10 to 20 times the variability in AUC. Hard to control

Systemic busulfan exposure, characterized by area under the plasma concentration curve, AUC = the delivered dose, is strongly associated with clinical outcome.

- AUC too high  $\Rightarrow$  High risks of severe toxicity.
- AUC too low  $\Rightarrow$  High risks of graft failure and disease recurrence.
- Earlier practice: busulfan orally, resulting in 10 to 20 times the variability in AUC. Hard to control
- Current practice: busulfan intravenous (IV), improving its bioavailability and delivered dosing accuracy

#### Cox proportional hazard model (details later):



#### Cox proportional hazard model (details later):



Andersson et al. (2002) estimated an optimal AUC range (950 to 1520  $\mu$ Mol-min)

#### Cox proportional hazard model (details later):



Andersson et al. (2002) estimated an optimal AUC range (950 to 1520  $\mu$ Mol-min)

Bartelink et al. (2016) reported an optimal AUC range (19100 to 21200  $\mu$ Mol-min) when treating children and young adults.

### **Personalizing Targeted IV Busulfan Dose**

Age and disease status at alloSCT (CR = Complete Remission, No CR = Active Disease) are strongly predictive of T = survival time.

### **Personalizing Targeted IV Busulfan Dose**

Age and disease status at alloSCT (CR = Complete Remission, No CR = Active Disease) are strongly predictive of T = survival time.

Can the optimal AUC interval be personalized using (Age, CR) to maximize  $E(T \mid Age, CR, AUC)$ ?

### **Personalizing Targeted IV Busulfan Dose**

Age and disease status at alloSCT (CR = Complete Remission, No CR = Active Disease) are strongly predictive of T = survival time.

Can the optimal AUC interval be personalized using (Age, CR) to maximize  $E(T \mid Age, CR, AUC)$ ?

**Historical Data:** 151 alloSCT patients who received a standard 4day preparative regimen of IV busulfan.

## **Challenges and Potential Benefits**

#### **Challenges in this Statistical Analysis**

# **Challenges and Potential Benefits**

#### **Challenges in this Statistical Analysis**

• Modeling  $p(T \mid Age, CR, AUC)$  robustly.

# **Challenges and Potential Benefits**

#### **Challenges in this Statistical Analysis**

- Modeling  $p(T \mid Age, CR, AUC)$  robustly.
- Identifying and characterizing possibly nonlinear [Age  $\times$  AUC ] or [CR  $\times$ AUC] or [Age  $\times$ CR  $\times$ AUC] interactive effects on *T*.
### **Challenges and Potential Benefits**

#### **Challenges in this Statistical Analysis**

- Modeling  $p(T \mid Age, CR, AUC)$  robustly.
- Identifying and characterizing possibly nonlinear [Age × AUC ] or [CR × AUC] or [Age × CR × AUC] interactive effects on T.
- The historical dataset has only 151 patients.

### **Challenges and Potential Benefits**

#### **Challenges in this Statistical Analysis**

- Modeling  $p(T \mid Age, CR, AUC)$  robustly.
- Identifying and characterizing possibly nonlinear [Age  $\times$  AUC ] or [CR  $\times$ AUC] or [Age  $\times$ CR  $\times$ AUC] interactive effects on *T*.
- The historical dataset has only 151 patients.

#### **Potential Payoff of this Statistical Analysis**

### **Challenges and Potential Benefits**

#### **Challenges in this Statistical Analysis**

- Modeling  $p(T \mid Age, CR, AUC)$  robustly.
- Identifying and characterizing possibly nonlinear [Age  $\times$  AUC ] or [CR  $\times$ AUC] or [Age  $\times$ CR  $\times$ AUC] interactive effects on *T*.
- The historical dataset has only 151 patients.

#### **Potential Payoff of this Statistical Analysis**

If optimal AUC intervals based on (Age, CR) can be estimated,  $E(T \mid Age, CR, AUC)$  can be increased for future patients by making this personalized IV busulfan dosing in alloSCT standard clinical practice.

Event time data, usually involving censoring.

- Survival time: T
- Censoring time: C
- Censoring indicator:  $\delta = I(T \le C)$
- Observed data:  $(Y, C, \delta)$ , where Y = min(T, C)

• Survival function: S(t) = Pr(T > t).

• Survival function: S(t) = Pr(T > t).

• **Density**: 
$$f(t) = -\frac{d}{dt}[1 - F(t)] = -\frac{d}{dt}S(t)$$
.

• Survival function: S(t) = Pr(T > t).

• **Density**: 
$$f(t) = -\frac{d}{dt}[1 - F(t)] = -\frac{d}{dt}S(t)$$
.

• Hazard function:  $\lambda(t) = \lim_{dt \to 0} \frac{Pr(t \le T < t + dt)}{dtS(t)} = \frac{f(t)}{S(t)}.$ 

• Survival function: S(t) = Pr(T > t).

• **Density**: 
$$f(t) = -\frac{d}{dt}[1 - F(t)] = -\frac{d}{dt}S(t)$$
.

- Hazard function:  $\lambda(t) = \lim_{dt \to 0} \frac{Pr(t \le T < t + dt)}{dtS(t)} = \frac{f(t)}{S(t)}.$
- Cumulative hazard function:  $\Lambda(t) = \int_{0}^{t} \lambda(u) du = -\log S(t).$

#### Kaplan–Meier estimator







 $t_i$ : time when at least one event happened  $d_i$ : the number of events (e.g., death) that happened at time  $t_i$  $n_i$ : the *individuals known to have survived* (have not yet had an event or been censored) up to time  $t_i$ 

• Cox Proportional Hazards (PH) model:

$$S_x(t) = S_0(t)^{\exp(x'\beta)}$$

• Cox Proportional Hazards (PH) model:

 $S_x(t) = S_0(t)^{\exp(x'\beta)}$ 

• In terms of hazards, this model reduces to

$$h_x(t) = h_0(t)\exp(x'\beta)$$

• Cox Proportional Hazards (PH) model:

 $S_x(t) = S_0(t)^{\exp(x'\beta)}$ 

• In terms of hazards, this model reduces to

$$h_x(t) = h_0(t)\exp(x'\beta)$$

Note then that for two individuals with covariates  $x_1$  and  $x_2$ , the ratio of hazard curves is constant, equal to  $\exp((x_1 - x_2)'\beta)$ , hence the name "proportional hazards."





Cox Proportional Hazards (PH) model is not appropriate!

• Accelerated Failure Time (AFT) model:

 $S_x(t) = S_0\{\exp(-x'\beta)t\}$ 

• Accelerated Failure Time (AFT) model:

 $S_x(t) = S_0\{\exp(-x'\beta)t\}$ 

• This is equivalent to a linear model for the log time-toevent *T*,

 $\log(T) = x'\beta + \epsilon$ , where  $p(\epsilon > \log t) = S_0(t)$ .

$$F(y \mid X) = \sum_{h=1}^{\infty} w_h N(y; \theta_h(X), \sigma^2)$$

$$F(y \mid X) = \sum_{h=1}^{\infty} w_h N(y; \theta_h(X), \sigma^2)$$

$$F(y \mid X) = \sum_{h=1}^{\infty} w_h N(y; \theta_h(X), \sigma^2)$$

• Include regression on covariates by assuming  $\theta_h(X) = X\beta_h$ .

$$F(y \mid X) = \sum_{h=1}^{\infty} w_h N(y; \theta_h(X), \sigma^2)$$

- Include regression on covariates by assuming  $\theta_h(X) = X\beta_h$ .
- A Gaussian process (GP) prior for  $\theta_h(X)$  gives the DDP-GP.  $\theta_h(X) \sim GP(\mu_h, C)$  with  $\mu_h(X_i; \beta_h) = X_i\beta_h$  for i = 1, ..., n and h = 1, 2, ..., with variance-covariance matrix

$$F(y \mid X) = \sum_{h=1}^{\infty} w_h N(y; \theta_h(X), \sigma^2)$$

- Include regression on covariates by assuming  $\theta_h(X) = X\beta_h$ .
- A Gaussian process (GP) prior for  $\theta_h(X)$  gives the DDP-GP.  $\theta_h(X) \sim GP(\mu_h, C)$  with  $\mu_h(X_i; \beta_h) = X_i\beta_h$  for i = 1, ..., n and h = 1, 2, ..., with variance-covariance matrix

$$C(\boldsymbol{Z}_i, \boldsymbol{Z}_\ell) = \sigma_0^2 \exp\left\{-\sum_{d=1}^D \frac{(Z_{id} - Z_{\ell d})^2}{\lambda_d^2}\right\} + \delta_{i\ell} J^2.$$

# **Data:** $Y_i = \log$ survival time, $X_i = (Age_i, CR_i, AUC_i)$ of patient i = 1, ..., 151

**Data:**  $Y_i = \log$  survival time,  $X_i = (Age_i, CR_i, AUC_i)$  of patient i = 1, ..., 151

**Distribution:**  $p(y_i | X_i, F) = F_{X_i}(y_i)$ 

**Data:**  $Y_i = \log$  survival time,  $X_i = (Age_i, CR_i, AUC_i)$  of patient i = 1, ..., 151

**Distribution:**  $p(y_i | X_i, F) = F_{X_i}(y_i)$ 

**Prior:**  $F_X \sim \text{DDP} - \text{GP}(\{\mu_h\}, C, \alpha, \{\beta_h\}, \{\lambda_d\}, \sigma_0^2, \sigma^2)$ 

**Data:**  $Y_i = \log$  survival time,  $X_i = (Age_i, CR_i, AUC_i)$  of patient i = 1, ..., 151

**Distribution:**  $p(y_i | X_i, F) = F_{X_i}(y_i)$ 

**Prior:**  $F_X \sim \text{DDP} - \text{GP}(\{\mu_h\}, C, \alpha, \{\beta_h\}, \{\lambda_d\}, \sigma_0^2, \sigma^2)$ 

For right-censored survival data  $D_n = \{Y_i, \delta_i, X_i\}_{i=1}^n$ , the likelihood function has the usual form

$$L(\theta \mid D_n) = \prod_{i=1}^n \{ f_{X_i}(Y_i \mid \theta) \}^{\delta_i} \{ 1 - F_{X_i}(Y_i \mid \theta) \}^{1 - \delta_i}$$

#### Estimated **Optimal Targeted Intervals** of IV Busulfan AUC Personalized For Given (CR Status, Age)

We define the predicted optimal IV busulfan targeted AUC for future patient n+1 with covariates X = (CR Status, Age, AUC) as

$$\widehat{AUC}_{n+1} = \operatorname{argmax}_{AUC} E(Y_{n+1} \mid \boldsymbol{X}, \mathcal{D}_n)$$

Since the laboratory error in evaluation of AUC is up to about 6%, the optimal AUC interval for future patient n + 1 is defined as

$$\begin{bmatrix} 0.9 \ \widehat{AUC}_{n+1}, & 1.1 \ \widehat{AUC}_{n+1} \end{bmatrix}$$

# Simulation Study: Setup

- Age: x<sub>1</sub> and AUC: x<sub>2</sub> were sampled with replacement from the actual ages and AUC values.
- CR:  $x_3 \sim \text{Bernoulli}(0.5)$ .
- Survival:  $T \sim LN(\mu(\mathbf{x}_i), \sigma_0^2)$ , where  $\sigma_0 = 0.4$ , and

$$\mu(\mathbf{x}_{i}) = 4 - 0.1x_{i,1} + 0.7x_{i,2} + 0.3x_{i,3} - 0.07x_{i,2}^{2} - 0.1x_{i,1}x_{i,2} + 0.2x_{i,2}x_{i,3} - 0.18x_{i,1}x_{i,2}x_{i,3}$$

• Two scenarios: n = 200 observations without censoring and n = 200 with 25% censoring.

# Simulation Study: Setup

- Age: x<sub>1</sub> and AUC: x<sub>2</sub> were sampled with replacement from the actual ages and AUC values.
- CR:  $x_3 \sim \text{Bernoulli}(0.5)$ .
- Survival:  $T \sim LN(\mu(\mathbf{x}_i), \sigma_0^2)$ , where  $\sigma_0 = 0.4$ , and

$$\mu(\mathbf{x}_{i}) = 4 - 0.1x_{i,1} + 0.7x_{i,2} + 0.3x_{i,3} - 0.07x_{i,2}^{2} - 0.1x_{i,1}x_{i,2} + 0.2x_{i,2}x_{i,3} - 0.18x_{i,1}x_{i,2}x_{i,3}$$

• Two scenarios: n = 200 observations without censoring and n = 200 with 25% censoring.

# Simulation Study: Setup

- Age: x<sub>1</sub> and AUC: x<sub>2</sub> were sampled with replacement from the actual ages and AUC values.
- CR:  $x_3 \sim \text{Bernoulli}(0.5)$ .
- Survival:  $T \sim LN(\mu(\mathbf{x}_i), \sigma_0^2)$ , where  $\sigma_0 = 0.4$ , and

$$\mu(\mathbf{x}_{i}) = 4 - 0.1x_{i,1} + 0.7x_{i,2} + 0.3x_{i,3} - 0.07x_{i,2}^{2} - 0.1x_{i,1}x_{i,2} + 0.2x_{i,2}x_{i,3} - 0.18x_{i,1}x_{i,2}x_{i,3}$$

• Two scenarios: n = 200 observations without censoring and n = 200 with 25% censoring.

#### Comparison

#### Comparison

• AFT regression models using either lognormal or Weibull distributions by assuming

#### Comparison

• AFT regression models using either lognormal or Weibull distributions by assuming

$$\log(Y_{i}) = \beta_{0} + \beta_{1}x_{i,1} + \beta_{2}x_{i,2} + \beta_{3}x_{i,3} + \beta_{4}x_{i,2}^{2} + \beta_{5}x_{i,1}x_{i,2} + \beta_{6}x_{i,2}x_{i,3} + \beta_{7}x_{i,1}x_{i,3} + \sigma\epsilon_{i}.$$

#### Comparison

• AFT regression models using either lognormal or Weibull distributions by assuming

$$\log(Y_{i}) = \beta_{0} + \beta_{1}x_{i,1} + \beta_{2}x_{i,2} + \beta_{3}x_{i,3} + \beta_{4}x_{i,2}^{2} + \beta_{5}x_{i,1}x_{i,2} + \beta_{6}x_{i,2}x_{i,3} + \beta_{7}x_{i,1}x_{i,3} + \sigma\epsilon_{i}.$$

- Two flexible semiparametric survival methods: model the baseline survival using
  - a Polya Trees (PT) prior (Hanson and Johnson, 2002)
#### Comparison

$$\log(Y_{i}) = \beta_{0} + \beta_{1}x_{i,1} + \beta_{2}x_{i,2} + \beta_{3}x_{i,3} + \beta_{4}x_{i,2}^{2} + \beta_{5}x_{i,1}x_{i,2} + \beta_{6}x_{i,2}x_{i,3} + \beta_{7}x_{i,1}x_{i,3} + \sigma\epsilon_{i}.$$

- Two flexible semiparametric survival methods: model the baseline survival using
  - a Polya Trees (PT) prior (Hanson and Johnson, 2002)
  - a transformed Bernstein polynomials (TBP) prior (Zhou and Hanson, 2018)

#### Comparison

$$\log(Y_{i}) = \beta_{0} + \beta_{1}x_{i,1} + \beta_{2}x_{i,2} + \beta_{3}x_{i,3} + \beta_{4}x_{i,2}^{2} + \beta_{5}x_{i,1}x_{i,2} + \beta_{6}x_{i,2}x_{i,3} + \beta_{7}x_{i,1}x_{i,3} + \sigma\epsilon_{i}.$$

- Two flexible semiparametric survival methods: model the baseline survival using
  - a Polya Trees (PT) prior (Hanson and Johnson, 2002)
  - a transformed Bernstein polynomials (TBP) prior (Zhou and Hanson, 2018)
- Two nonparametric survival methods

#### Comparison

$$\log(Y_{i}) = \beta_{0} + \beta_{1}x_{i,1} + \beta_{2}x_{i,2} + \beta_{3}x_{i,3} + \beta_{4}x_{i,2}^{2} + \beta_{5}x_{i,1}x_{i,2} + \beta_{6}x_{i,2}x_{i,3} + \beta_{7}x_{i,1}x_{i,3} + \sigma\epsilon_{i}.$$

- Two flexible semiparametric survival methods: model the baseline survival using
  - a Polya Trees (PT) prior (Hanson and Johnson, 2002)
  - a transformed Bernstein polynomials (TBP) prior (Zhou and Hanson, 2018)
- Two nonparametric survival methods
  - random forests (RF) (Ishwaran et al., 2008)

#### Comparison

$$\log(Y_{i}) = \beta_{0} + \beta_{1}x_{i,1} + \beta_{2}x_{i,2} + \beta_{3}x_{i,3} + \beta_{4}x_{i,2}^{2} + \beta_{5}x_{i,1}x_{i,2} + \beta_{6}x_{i,2}x_{i,3} + \beta_{7}x_{i,1}x_{i,3} + \sigma\epsilon_{i}.$$

- Two flexible semiparametric survival methods: model the baseline survival using
  - a Polya Trees (PT) prior (Hanson and Johnson, 2002)
  - a transformed Bernstein polynomials (TBP) prior (Zhou and Hanson, 2018)
- Two nonparametric survival methods
  - random forests (RF) (Ishwaran et al., 2008)
  - Bayesian additive regression trees (BART) (Sparapani et al., 2016)

#### n=200 without censoring

n=200 with 25% censoring



#### n=200 without censoring



CR=0

CR=1

n=200 with 25% censoring



CR=0

CR=1

#### **Kaplan Meier Plots**



Time

AUC=5









• The optimal targeted busulfan dose interval goes down with Age.



- The optimal targeted busulfan dose interval goes down with Age.
- For Age  $\leq$  30, the optimal targeted IV busulfan dose intervals are identical for patients in CR or with active disease (no CR).



- The optimal targeted busulfan dose interval goes down with Age.
- For Age  $\leq$  30, the optimal targeted IV busulfan dose intervals are identical for patients in CR or with active disease (no CR).
- For Age > 30, the optimal targeted dose intervals for [CR = No] are well below the intervals for [CR = Yes], with complete separation for Age > 55.

• This precision IV busulfan dosing may be applied worldwide in alloSCT to improve survival.

- This precision IV busulfan dosing may be applied worldwide in alloSCT to improve survival.
- The DDP-GP model is a tool for robust Bayesian nonparametric survival regression analysis that may be applied widely.

- This precision IV busulfan dosing may be applied worldwide in alloSCT to improve survival.
- The DDP-GP model is a tool for robust Bayesian nonparametric survival regression analysis that may be applied widely.

The R package DDPGPSurv can be downloaded from https://cran.r-project.org/web/packages/DDPGPSurv/index.html